

The funnel plot appears asymmetric, and there is evidence of bias using the Egger (weighted regression) method (P for bias 0.007) but not using the Begg (rank correlation method). This is compatible with a greater statistical power of the regression test, as discussed in Chapter 11. The horizontal line in the funnel plot indicates the fixed-effects summary estimate (using inverse-variance weighting), while the sloping lines indicate the expected 95% confidence intervals for a given standard error, assuming no heterogeneity between studies.

### Meta-regression

If evidence is found of heterogeneity in the effect of treatment between studies, then meta-regression can be used to analyse associations between treatment effect and study characteristics. Meta-regression can be done in Stata by using the `metareg` command.<sup>19</sup>

#### Example 3: trials of BCG vaccine against tuberculosis

The following table is based on a meta-analysis by Colditz *et al.*<sup>20</sup> which examined the efficacy of BCG vaccine against tuberculosis.

Table 18.3

Trial	Trial name	Authors	Start year	Latitude*	Intervention group		Control group	
					TB cases	Total cases	TB cases	Total cases
1	Canada	Ferguson & Simes	1933	55	6	306	29	303
2	Northern USA	Aronson	1935	52	4	123	11	139
3	Northern USA	Stein & Aronson	1935	52	180	1541	372	1451
4	Chicago	Rosenthal <i>et al.</i>	1937	42	17	1716	65	1665
5	Chicago	Rosenthal <i>et al.</i>	1941	42	3	231	11	220
6	Georgia (School)	Comstock & Webster	1947	33	5	2498	3	2341
7	Puerto Rico	Comstock <i>et al.</i>	1949	18	186	50634	141	27338
8	UK	Hart & Sutherland	1950	53	62	13598	248	12867
9	Madanapalle	Frimont-Moller <i>et al.</i>	1950	13	33	5069	47	5808
10	Georgia (Community)	Comstock <i>et al.</i>	1950	33	27	16913	29	17854
11	Haiti	Vandeviere <i>et al.</i>	1965	18	8	2545	10	629
12	South Africa	Coetzee & Barjak	1965	27	29	7499	45	7277
13	Madras	TB prevention trial	1968	13	505	88391	499	88391

\* Expressed in degrees from equator.

The data were saved in Stata dataset `bcgtrial.dta`.

### describe

```
Contains data from bcgtrial.dta
obs:      13
vars:     9
size:     754 (99.9% of memory free)
```

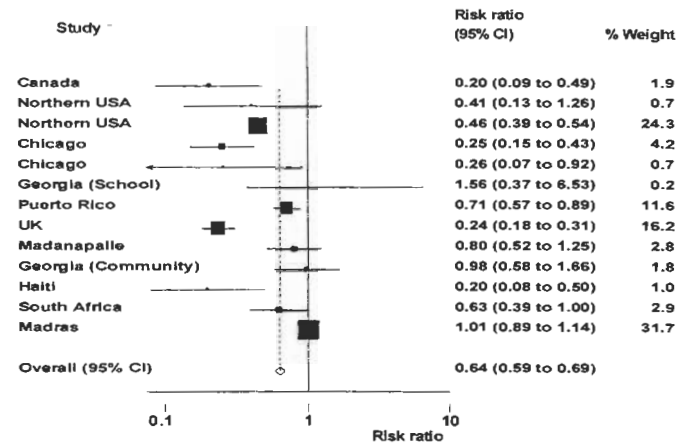
```
-----+-----
1. trial      byte      %8.0g
2. trialnam   str19     %19s
3. authors   str19     %19s
4. startyr   int        %8.0g
5. latitude  byte      %8.0g
6. cases1    int        %8.0g
7. tot1      long      %12.0g
8. cases0    int        %8.0g
9. tot0      long      %12.0g
-----+-----
```

Sorted by: trial

Scientists had been aware of discordance between the results of these trials since the 1950s. The clear heterogeneity in the protective effect of BCG between trials can be seen in the forest plot (we analyse this study using risk ratios):

```
gen h1=tot1-cases1
gen h0=tot0-cases0

metan cases1 h1 cases0 h0, xlab(.1,1,10)
label(namevar=trialnam)
```



To use the `metareg` command, we need to derive the treatment effect estimate (in this case log risk ratio) and its standard error, for each study.

```
generate logrr=log((cases1/tot1)/(cases0/tot0))
generate selogrr=sqrt((1/cases1)-(1/tot1)+(1/cases0)-(1/tot0))
```

In their meta-analysis, Colditz *et al.* noted the strong evidence for heterogeneity between studies, and concluded that a random-effects meta-analysis was appropriate:

```
meta logrr selogrr, eform
Meta-analysis (exponential form)
```

Method	Pooled			Asymptotic		No. of studies
	Est	Lower	Upper	z_value	p_value	
Fixed	0.650	0.601	0.704	-10.625	0.000	13
Random	0.490	0.345	0.695	-3.995	0.000	

Test for heterogeneity:  $Q = 152.233$  on 12 degrees of freedom ( $p = 0.000$ )  
 Moment-based estimate of between studies variance = 0.309

(The different weight of studies under the fixed and random effects assumption is discussed in Chapter 2).

The authors then examined possible explanations for the clear differences in the effect of BCG between studies. The earlier studies may have produced different results than later ones. The latitude at which the studies were conducted may also be associated with the effect of BCG. As discussed by Fine,<sup>21</sup> the possibility that BCG might provide greater protection at higher latitudes was first recognised by Palmer and Long,<sup>22</sup> who suggested that this trend might result from exposure to certain environmental mycobacteria, more common in warmer regions, which impart protection against tuberculosis.

To use `metareg`, we provide a list of variables, the first of which is the treatment effect (here, the log risk ratio) and the rest of which are (one or more) study characteristics (covariates) hypothesized to be associated with the treatment effect. In addition, the standard error or variance of the treatment effect must be provided, using the `wsse` (within-study standard error) or `wsvar` (within-study variance) option. It is also possible to specify the method for estimating the between-study variance: here we use the default; restricted maximum-likelihood (`reml`). To look for an association with start year and latitude:

```
metareg logrr startyr latitude, wsse(selogrr)
```

```
Iteration 1: tau^2 = 0
Iteration 2: tau^2 = .02189942
:
:
Iteration 9: tau^2 = .1361904
Iteration 10: tau^2 = .13635174
```

```
Meta-analysis regression      No of studies = 13
                             tau^2 method   reml
                             tau^2 estimate = .1364
```

Successive values of tau^2 differ by less than 10^-4 : convergence achieved

	Coef.	Std. Err.	z	P> z	[95% Conf.Interval]
startyr	-.004966	.0162811	-0.305	0.760	-.0368763 .0269444
latitude	-.0270477	.0118195	-2.288	0.022	-.0502135 -.0038819
_cons	9.890987	32.02516	0.309	0.757	-52.87717 72.65914

The regression coefficients are the estimated increase in the log risk ratio per unit increase in the covariate. So in the example the log risk ratio is estimated to decrease by 0.027 per unit increase in the latitude at which the study is conducted. The estimated between-study variance has been reduced from 0.31 (see output from the `meta` command) to 0.14. While there is strong evidence for an association between latitude and the effect of BCG, there is no evidence for an association with the year the study started. The estimated treatment effect given particular values of the covariates may be derived from the regression equation. For example, for a trial beginning in 1950, at latitude 50°, the estimated log risk ratio is given by:

$$\text{Log risk ratio} = 9.891 - 0.00497 \times 1950 - 0.0270 \times 50 = -1.1505$$

which corresponds to a risk ratio of  $\exp(-1.1505) = 0.316$

The use of meta-regression in explaining heterogeneity and identifying sources of bias in meta-analysis is discussed further in Chapters 8–11.

- 1 Yusuf S, Collins R, Peto R, *et al.* Intravenous and intracoronary fibrinolytic therapy in acute myocardial infarction: overview of results on mortality, reinfarction and side-effects from 33 randomized controlled trials. *Eur Heart J* 1985;6:556–85.
- 2 Gruppo Italiano per lo Studio della Streptochinasi nell'Infarto Miocardico (GISSI). Effectiveness of intravenous thrombolytic treatment in acute myocardial infarction. *Lancet* 1986;1:397–402.
- 3 ISIS-2 (Second International Study of Infarct Survival) Collaborative Group. Randomised trial of intravenous streptokinase, oral aspirin, both, or neither among 17,187 cases of suspected acute myocardial infarction: ISIS-2. *Lancet* 1988;2:349–60.
- 4 Bradburn MJ, Deeks JJ, Altman DG. `sbc24: metan` – an alternative meta-analysis command. *Stata Tech Bull* 1998;44:15.
- 5 Sharp S, Sterne J. `sbc16: Meta-analysis`. *Stata Tech Bull* 1997;38:9–14.
- 6 Sharp S, Sterne J. `sbc16.1: New syntax and output for the meta-analysis command`. *Stata Tech Bull* 1998;42:6–8.
- 7 Sharp S, Sterne J. `sbc16.2: Corrections to the meta-analysis command`. *Stata Tech Bull* 1998;43:15.
- 8 Teo KK, Yusuf S, Collins R, Held PH, Peto R. Effects of intravenous magnesium in suspected acute myocardial infarction: overview of randomised trials. *BMJ* 1991;303:1499–503.
- 9 ISIS-4 (Fourth International Study of Infarct Survival) Collaborative Group. ISIS-4: a randomised factorial trial assessing early oral captopril, oral mononitrate, and intravenous magnesium sulphate in 58,050 patients with suspected acute myocardial infarction. *Lancet* 1995;345:669–85.
- 10 Egger M, Smith GD. Misleading meta-analysis. Lessons from an “effective, safe, simple” intervention that wasn’t. *BMJ* 1995;310:752–4.
- 11 Egger M, Smith GD, Schneider M, Minder C. Bias in meta-analysis detected by a simple, graphical test. *BMJ* 1997;315:629–34.
- 12 Sterne J. `sbc22: Cumulative meta analysis`. *Stata Tech Bull* 1998;42:13–16.
- 13 Lau J, Antman EM, Jimenez-Silva J, Kupelnick B, Mosteller F, Chalmers TC. Cumulative meta-analysis of therapeutic trials for myocardial infarction. *N Engl J Med* 1992;327:248–54.
- 14 Antman EM, Lau J, Kupelnick B, Mosteller F, Chalmers TC. A comparison of results of meta-analyses of randomized control trials and recommendations of clinical experts’ Treatments for myocardial infarction. *JAMA* 1992;268:240–8.
- 15 Tobias A. `sbc26: Assessing the influence of a single study in meta-analysis`. *Stata Tech Bull* 1999;47:15–17.

- 16 Steichen T. `sbc19: Tests for publication bias in meta-analysis`. *Stata Tech Bull* 1998;41:9–15.
- 17 Steichen T, Egger M, Sterne J. `sbc19.1: Tests for publication bias in meta-analysis`. *Stata Tech Bull* 1998;44:3–4.
- 18 Begg CB, Mazumdar M. Operating characteristics of a rank correlation test for publication bias. *Biometrics* 1994;50:1088–101.
- 19 Sharp S. `sbc23: Meta-analysis regression`. *Stata Tech Bull* 1998;42:16–24.
- 20 Colditz GA, Brewer TJ, Berkey CS, *et al.* Efficacy of BCG vaccine in the prevention of tuberculosis. Meta-analysis of the published literature. *JAMA* 1994;271:698–702.
- 21 Fine PEM. Variation in protection by BCG: implications of and for heterologous immunity. *Lancet* 1995;346:1339–45.
- 22 Palmer CE, Long MW. Effects of infection with atypical mycobacteria on BCG vaccination and tuberculosis. *Am Rev Respir Dis* 1966;94:553–68.