Coupled climate models are sophisticated tools designed to simulate the Earth climate system and the complex interactions between its components. Currently, more than a dozen centers around the world develop climate models to enhance our understanding of climate and climate change and to support the activities of the Intergovernmental Panel on Climate Change (IPCC). However, climate models are not perfect. Our theoretical understanding of climate is still incomplete, and certain simplifying assumptions are unavoidable when building these models. This introduces biases into their simulations, which sometimes are surprisingly difficult to correct. Model imperfections have attracted criticism, with some arguing that model-based projections of climate are too unreliable to serve as a basis for public policy. In particular, early attempts at coupled modeling in the 1980s resulted in relatively crude representations of climate. Since then, however, we have refined our theoretical understanding of climate, improved the physical basis for climate modeling, increased the number and quality of observations, and multiplied our computational capabilities. Against the background of these developments, one may ask how much climate models have improved and how much we can trust the latest coupled model generation.

The goal of this study is to objectively quantify the agreement between model and observations using a single quantity derived from a broad group of variables, which is then applied to gauge several generations of coupled climate models. This approach is new, since previous model intercomparison studies either focused on specific processes, avoided making quantitative performance statements, or considered a rather narrow range of models.

Several important issues complicate the model validation process. First, identifying model errors is difficult because of the complex and sometimes poorly understood nature of climate itself, making it difficult to decide which of the many aspects of climate are important for a good simulation. Second, climate models must be compared against present (e.g., 1979–99) or past climate, since verifying observations for future climate are unavailable. Present climate, however, is not an independent dataset since it has already been used for the model development. On the other hand, information about past climate carries large inherent uncertainties, complicating the validation process of past climate simulations. Third, there is a lack of reliable and consistent observations for present climate, and some climate processes occur at temporal or spatial scales that are either unobservable or unresolvable. Finally, good model performance evaluated from the present climate does not necessarily guarantee reliable predictions of future climate. Despite these difficulties and limitations, model agreement with observations of today’s climate is the only way to assign model confidence, with the underlying assumption that a model that accurately describes present climate will make a better projection of the future.

Considering the above complications, it is clear that there is no single “ideal way” to characterize and compare model performances. Most previous model validation studies used conventional statistics to measure the similarity between observed and modeled data. For example, studies by Taylor et al. (2001) and Boer and Lambert (2001) characterized model performance from correlation, root-mean-square (RMS) error, and variance ratio. Both studies found similar ways to combine these three statistics.
in a single diagram, resulting in nice graphical visualizations of model performance. This approach, however, is only practical for a small number of models and/or climate quantities. In addition, Taylor's widely used approach requires centered RMS errors with the mean bias removed. We, however, consider the mean bias as an important component of model error. In a 2004 article, Murphy et al. introduced a Climate Prediction Index (CPI), which measures the reliability of a model based on the composite mean-square errors of a broad range of climate variables. More recently, Min and Hense (2006) introduced a Bayesian approach into model evaluation, where skill is measured in terms of a likelihood ratio of a model with respect to some reference.

**THREE GENERATIONS OF MODEL DATA.**

This study includes model output from three different climate model intercomparison projects (CMIP): CMIP-1 (Meehl et al. 2000), the first project of its kind organized in the mid-1990s; the follow-up project CMIP-2 (Covey et al. 2003, Meehl et al. 2005); and CMIP-3 (PCMDI 2007) (aka, IPCC-AR4), representing today's state of the art in climate modeling. The CMIP-3 data were taken from the "climate of the twentieth century" (20C3M) (hereafter simply "present-day") and the "preindustrial control" (PICNTRL) (hereafter simply "preindustrial") experiments. These simulations were driven by a rather realistic set of external forcings, which included the known or estimated history of a range of natural and anthropogenic sources, such as variations in solar output, volcanic activity, trace gases, and sulfate aerosols. The exact formulation of these forcings varied from model to model, with potential implications for model performance. In contrast, the CMIP-1 and CMIP-2 model output was derived from long control runs, in which the forcings were held constant in time. These forcings were only approximately representative for present climate.

**MEASURE OF MODEL PERFORMANCE.** As outlined before, there are many different ways to measure and depict model performance. Given the extra challenge of this study to evaluate and depict a large number of models and climate variables, we decided to design our own measure. Our strategy was to calculate a single performance index, which can be easily depicted, and which consists of the aggregated errors in simulating the observed climatological mean states of many different climate variables. We focused on validating the time-mean state of climate, since this is the most fundamental and best-observed aspect of climate, and because of restrictions imposed by available model data in calculating higher moments of climate (most CMIP-1 fields are archived as climatological means, prohibiting the derivation of temporal variability). This concept is somewhat similar to the CPI performance measure introduced by Murphy et al. (2004), but in contrast to the present study, Murphy et al. calculated the CPI from a range of rather closely related models.

Our choice of climate variables, which is shown in Table 1, was dictated by the data available from the models. In most cases, we were able to validate the model data against true observation-based data, but for a few variables of the free atmosphere, the usage of reanalyses as validation data was unavoidable. In terms of the specific uncertainties associated with each of those validating datasets, separate analysis showed that the data can be considered as good approximations to the real state of present climate for the purpose of model validation.

We obtained the model performance index by first calculating multiyear annual mean climatologies from global gridded fields of models and validating data. The base period for the observations was 1979–99, covering most of the well-observed post-1979 satellite period. For some observations, fewer years were used if data over the entire period were not available. For the CMIP-1 models, long-term climatologies of the control run for Northern Hemisphere winter (December, January, February) and summer (June, July, August) conditions were downloaded from the archives and averaged to annual mean climatologies. The CMIP-2 climatologies were calculated by averaging the annual mean data of the control run over the years 61–80. The CMIP-3 present-day climatologies were formed using the same base period as for the observations, and the preindustrial climatologies were taken from the last 20 simulation years of the corresponding control run. For any given model, only one member integration was included. In the rare case that a climate variable was not provided by a specific model, we replaced the unknown error by the mean error over the remaining models of the corresponding model generation. One model (BCC-CM1 from CMIP-3) was excluded because it only provided a small subset of variables needed for this study.

In determining the model performance index, we first calculated for each model and variable a normalized error variance $e^2$ by squaring the grid-point differences between simulated (interpolated to the observational grid) and observed climate, normalizing
TABLE I. Climate variables and corresponding validation data. Variables listed as “zonal mean” are latitude–height distributions of zonal averages on twelve atmospheric pressure level between 1000 and 100 hPa. Those listed as “ocean,” “land,” or “global” are single-level fields over the respective regions. The variable “net surface heat flux” represents the sum of six quantities: incoming and outgoing shortwave radiation, incoming and outgoing longwave radiation, and latent and sensible heat fluxes. Period indicates years used to calculate observational climatologies.

<table>
<thead>
<tr>
<th>Variable</th>
<th>Domain</th>
<th>Validation data</th>
<th>Period</th>
</tr>
</thead>
<tbody>
<tr>
<td>Sea level pressure</td>
<td>ocean</td>
<td>ICOADS (Woodruff et al. 1987)</td>
<td>1979–99</td>
</tr>
<tr>
<td>Air temperature</td>
<td>zonal</td>
<td>ERA-40 (Simmons and Gibson 2000)</td>
<td>1979–99</td>
</tr>
<tr>
<td>Zonal wind stress</td>
<td>ocean</td>
<td>ICOADS (Woodruff et al. 1987)</td>
<td>1979–99</td>
</tr>
<tr>
<td>Meridional wind stress</td>
<td>ocean</td>
<td>ICOADS (Woodruff et al. 1987)</td>
<td>1979–99</td>
</tr>
<tr>
<td>2-m air temperature</td>
<td>global</td>
<td>CRU (Jones et al. 1999)</td>
<td>1979–99</td>
</tr>
<tr>
<td>Zonal wind</td>
<td>zonal</td>
<td>ERA-40 (Simmons and Gibson 2000)</td>
<td>1979–99</td>
</tr>
<tr>
<td>Meridional wind</td>
<td>zonal</td>
<td>ERA-40 (Simmons and Gibson 2000)</td>
<td>1979–99</td>
</tr>
<tr>
<td>Precipitation</td>
<td>global</td>
<td>CMAP (Xie and Arkin 1998)</td>
<td>1979–99</td>
</tr>
<tr>
<td>Specific humidity</td>
<td>zonal</td>
<td>ERA-40 (Simmons and Gibson 2000)</td>
<td>1979–99</td>
</tr>
<tr>
<td>Sea surface temperature</td>
<td>ocean</td>
<td>GISSST (Parker et al. 1995)</td>
<td>1979–99</td>
</tr>
<tr>
<td>Sea ice fraction</td>
<td>ocean</td>
<td>GISSST (Parker et al. 1995)</td>
<td>1979–99</td>
</tr>
<tr>
<td>Sea surface salinity</td>
<td>ocean</td>
<td>NODC (Levitus et al. 1998)</td>
<td>variable</td>
</tr>
</tbody>
</table>

The final model performance index was formed by taking the mean over all climate variables (Table 1) and one model using equal weights,

$$I_m^2 = \frac{\overline{I}_{vm}^2}{\overline{e}_{vm}^2},$$

where the overbar indicates averaging. The reference ensemble was the present-day CMIP-3 experiment. The final step combines the errors from different climate variables into one index. We justify this step by normalizing the individual error components prior to taking averages [Eqs. (1) and (2)]. This guarantees that each component varies evenly around one and has roughly the same variance. In this sense, the individual $I_{vm}^2$ values can be understood as rankings with respect to individual climate variables, and the final index is the mean over all ranks. Note that a very similar approach has been taken by Murphy et al. (2004).

RESULTS. The outcome of the comparison of the 57 models in terms of the performance index $I^2$ is illustrated in the top three rows of Fig. 1. The $I^2$ index varies around one, with values greater than one for underperforming models and values less than one...
for more accurate models. Since $I^2$ is an indicator of model performance relative to the mean over the present-day CMIP-3 ensemble, we used a logarithmic scale to display the index. The results indicate large differences from model to model in terms of their ability to match the observations of today’s climate. Further, the results clearly demonstrate a continuous improvement in model performance from the early CMIP-1 to the latest CMIP-3 generation. To our knowledge, this is the first systematic attempt to compare the performance of entire generations of climate models by exploring their ability to simulate present climate. Figure 1 also shows that the realism of the best models approaches that of atmospheric reanalysis (indicated by the green circle), but the models achieve this without being constrained by real observations.

We also obtained quantitative estimates of the robustness of the $I^2$ values by validating the models against a large synthetic ensemble of observational climatologies and by calculating the range of $I^2$ values encompassed by the 5th and 95th percentiles. The synthetic ensemble was produced by selecting the years included in each climatology using bootstrapping (i.e., random selection with replacement). To the extent that the circles in Fig. 1 overlap, it is not possible to distinguish the performance of the corresponding models in a way that is statistically significant.

**ROLE OF FORCINGS.** Given the more realistic forcing used for the present-day CMIP-3 simulations, the superior outcome of the corresponding models is perhaps not too surprising. One might ask how important realistic forcing was in producing such good simulations. To this end, we included the preindustrial CMIP-3 simulations in our comparison. Both the present-day and the preindustrial simulations were conducted with identical models. The only difference was the forcing used to drive the simulations, which was similar to preindustrial conditions for the preindustrial experiments and similar to present-day conditions for the present-day experiments.

The outcome of validating the preindustrial experiment against current climate is shown in the bottom row of Fig. 1. As expected, the $I^2$ values are now larger than for the present-day simulations, indicating poorer performance. However, the mean difference between the two CMIP-3 simulations, which was due only to different forcings, is much smaller than that between CMIP-3 and the previous two model generations. The latter difference was due to different models and forcings combined. We conclude that the superior performance of the CMIP-3 models is mostly related to drastic model improvements, and that the forcings used to drive these models play a more subtle role.

Two developments—more realistic parameterizations and finer resolutions—are likely to be most
Also note that CMIP-3 always performs better than CMIP-1, and almost always better than CMIP-2, even when only one variable is included. These results indicate that $I^2$, when used to compare entire model generations, is robust with respect to the number and choice of selected variables.

**VALUE OF THE MULTIMODEL MEAN.** We also investigated the performance of the multimodel means (black circles in Fig. 1), which are formed by averaging across the simulations of all models of one model generation and using equal weights. Notably, the multimodel mean usually outperforms any single model, and the CMIP-3 multimodel mean performs nearly as well as the reanalysis. Such performance improvement is consistent with earlier findings by Lambert and Boer (2001), Taylor et al. (2004), and Randall et al. (2007) regarding CMIP-1, AMIP-2, and CMIP-3 model output, respectively.

The use of multimodel ensembles is common practice in weather and short-term climate forecasting, and it is starting to become important for long-term climate change predictions. For example, many climate change estimates of the recently released global warming report of the IPCC are based on the multimodel simulations from the CMIP-3 ensemble. The report dealt with the problem of inconsistent predictions, resulting from the use of different models, by simply taking the average of all models as the best estimate for future climate change. Our results indicate that multimodel ensembles are a legitimate and effective means to improve the outcome of climate simulations. As yet, it is not exactly clear why the multimodel mean is better than any individual model.

**SENSITIVITY OF THE INDEX.** We now address the question of how sensitive our results are with respect to our particular choice of variables. We used bootstrapping to investigate how $I^2$—averaged individually over the four model groups—varies with an increasing number $v$ of variables. For any given $v$, we calculated $I^2$ many times, every time using different randomly chosen variable combinations taken from Table 1. As shown in Fig. 2, the spread of outcomes decreases with increasing number of variables. When six or more variables are used to calculate $I^2$, the average performances of the three model generations are well separated from each other—indeed from the exact choice of variables. Only the two CMIP-3 experiments cannot be distinguished from each other, even for a very large number of variables.

Fig. 2. Spread of $I^2$ values (lowest to highest) for an increasing number of randomly chosen variables $v$. Shown are index values averaged individually over the four model groups (corresponding to the grey circles in Fig. 1). In order to avoid nonunity results for 20C3M, all values were normalized by the mean $I^2$ over all three model generations, and not by the mean of the 20C3M group alone (as in Fig. 1, see Eq. 2).

Fig. 3. Fraction of flux-adjusted models among the three model generations.
model. One possible explanation is that the model solutions scatter more or less evenly about the truth (unless the errors are systematic), and the errors behave like random noise that can be efficiently removed by averaging. Such noise arises from internal climate variability, and probably to a much larger extent from uncertainties in the formulation of models.

**ROLE OF FLUX CORRECTION.** When discussing coupled model performances, one must take into account that earlier models are generally flux corrected, whereas most modern models do not require such corrections (Fig. 3). Flux correction, or adding artificial terms of heat, momentum, and freshwater at the air–sea interface, prevents models from drifting to unrealistic climate states when integrating over long periods of time. The drift, which occurs even under unforced conditions, is the result of small flux imbalances between ocean and atmosphere. The effects of these imbalances accumulate over time and tend to modify the mean temperature and/or salinity structure of the ocean. The technique of flux correction attracts concern because of its inherently nonphysical nature. The artificial corrections make simulations at the ocean surface more realistic, but only for artificial reasons. This is demonstrated by the increase in systematic biases (defined as the multimodel mean minus the observations) in sea surface temperatures from the mostly flux-corrected CMIP-1 models to the generally uncorrected CMIP-3 models (Fig. 4a). Because sea surface temperatures exert an important control on the exchange of properties across the air–sea interface, corresponding errors readily propagate to other climate fields. This can be seen in Fig. 4b, which shows that biases in ocean temperatures tend to be accompanied by same-signed temperature biases in the free troposphere. On the other hand, the reduction of strong lower stratospheric cold biases in the CMIP-3 models indicates considerable model improvements. These cold biases are likely related to the low vertical and horizontal resolution of former model generations and to the lack of parameterizations for small-scale gravity waves, which break, deposit momentum, and warm the middle atmosphere over the high latitudes. Modern models use appropriate parameterizations to replace the missing momentum deposition.

**CONCLUSION.** Using a composite measure of model performance, we objectively determined the ability of three generations of models to simulate present-day mean climate. Current models are certainly not perfect, but we found that they are much more realistic than their predecessors. This is mostly related to the enormous progress in model development that took place over the last decade, which is partly due to more sophisticated model parameterizations, but also to the general increase in computational resources, which allows for more thorough model testing and higher model resolut-
tion. Most of the current models not only perform better, they are also no longer flux corrected. Both improved performance and more physical formulation suggest that an increasing level of confidence can be placed in model-based predictions of climate. This, however, is only true to the extent that the performance of a model in simulating present mean climate is related to the ability to make reliable forecasts of long-term trends. It is hoped that these advancements will enhance the public credibility of model predictions and help to justify the development of even better models.

Given the many issues that complicate model validation, it is perhaps not too surprising that the present study has some limitations. First, we note the caveat that we were only concerned with the time-mean state of climate. Higher moments of climate, such as temporal variability, are probably equally as important for model performance, but we were unable to investigate these. Another critical point is the calculation of the performance index. For example, it is unclear how important climate variability is compared to the mean climate, exactly which is the optimum selection of climate variables, and how accurate the used validation data are. Another complicating issue is that error information contained in the selected climate variables is partly redundant. Clearly, more work is required to answer the above questions, and it is hoped that the present study will stimulate further research in the design of more robust metrics. For example, a future improved version of the index should consider possible redundancies and assign appropriate weights to errors from different climate variables. However, we do not think that our specific choices in this study affect our overall conclusion that there has been a measurable and impressive improvement in climate model performance over the past decade.

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FOR FURTHER READING


Jones, R., 2005: Senate hearing demonstrates wide disagreement about climate change. FYI Number 142, American Institute of Physics. [Available online at www.aip.org/fyi/2005/142.html.]


AMS 2008 Public–Private Partnership Forum
(Formerly the Corporate Forum)
University of California—Washington Center
Washington, D.C.
22–23 April 2008

SESSION TOPICS

Congressional Legislation—Authorization Committees
House and Senate staff will provide an outlook on pending legislation and discuss issues associated with programs and initiatives that may provide opportunities for AMS members.

Congressional Legislation—Appropriations Committees
House and Senate staff will provide an outlook on Federal budget appropriations for NASA, NOAA, NIST, and NSF for weather and climate related activities.

Agency Initiatives, Plans, and Opportunities
Senior staff from NOAA, NASA, DHS, and DOE will look ahead and provide updates on current meteorological, climatological, and oceanographic programs and provide insights on new science initiatives and directions.

Responding to Emerging Climate Change Policies
Climate change has resulted in a diverse response by policy makers, industry, and the media. This session will discuss the weather and climate community’s role in developing and implementing new climate services and products aimed at responding and supporting policies such as carbon trading, clean energy, and other climate change risk mitigation strategies.

Establishing Community Priorities—Defining, Prioritizing, and Advocating Effectively
Congress has repeatedly asked the weather and climate enterprise to clearly articulate our needs. In response, the AMS is preparing a statement on priorities. This session will provide an update on progress being made on this critically important topic.

PURPOSE
To provide an opportunity for weather and climate enterprise stakeholders to meet with senior Federal agency officials and Congressional staff to hear about the status of current programs, learn about new initiatives; discuss issues of interest to the enterprise, and identity business opportunities.

WHO SHOULD ATTEND
All AMS Members are encouraged to attend, particularly AMS Corporate Members, private sector executives, federal agency and laboratory managers and scientists, and university faculty.

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