LGN 5822 - Biometrical Genetics

## L10 – Latin Square

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## Introduction

 In some situations, there may be more than one source of heterogeneity among experimental units  The Latin Square Design is an appropriate design for environments heterogeneous experiments

In the Latin Square Design, in addition to the principles of repetition and randomization, the principle of local control is used twice to control the effect of two factors

#### More control of experimental units

## Introduction

 To control this variability, it is necessary to divide the experimental units into homogeneous blocks of experimental units in relation to each controlled factor

## Design Characterization

 If there are two gradients, in perpendicular directions, blocking factors can be used to simultaneously control both sources of variation

- Physical gradients in the field plot
- Other (orthogonal) experimental sources of variation
- Treatments of interest

When there are two main sources of variation that can be controlled

## Example

• Field trials in which the experimental error has two fertility gradients running perpendicular each other



Gradient 1

### Experimental Design Layout



В	O D	Δ _	் c
C c	О <sub>А</sub>	△в	Ο,
D	Ов	∧ <sub>c</sub>	
	⊖ <sub>c</sub>	∆ <sub>d</sub>	Ċ,

### Experimental Design Layout

 Consider a competition experiment with 4 sugarcane varieties in which the experimental area presents a soil fertility gradient in two directions.

Row 1	т1	Т2	Т3	Т4
Row 2	Т2	тз	Т4	T1
Row 3	тз	T4	T1	T2
Row 4	Т4	TI	T2	тз

Column 1 Column 2 Column 3 Column 4

## Description of the Design

- With the Latin Square design you are able to control variation in two directions
- Treatments are arranged in rows and columns
- Each row contains every treatment
- Each column contains every treatment
- The total number of experimental units is thus  $t^2$  (t t is the number of treatments)
  - This is a square design

## Description of the Design

• The number of treatments is equal to the number of repetitions

- This design is advisable when the number of treatments varies between 3 and 10
- But, for 3 and 4 treatments, only when the experiment can be repeated in several Latin squares
- It has more efficient local control than the randomized block design (horizontal and vertical control)

### Casualization in the Latin Square Design

 The treatments are distributed within the rows, so that each column also contains all the treatments

### Casualization in the Latin Square Design

 The treatments are distributed within the rows, so that each column also contains all the treatments

	Colunas						
Linhas	1	2	3	4	5		
1	Α	В	С	D	E		
2	Е	Α	В	С	D		
3	D	E	Α	В	С		
4	С	D	E	Α	В		
5	В	С	D	E	Α		

### Casualization in the Latin Square Design

Then, rows are then randomly distributed among themselves, and then the columns

Ε	Α	В	С	D
С	D	Ε	Α	В
В	С	D	Е	Α
Α	В	С	D	Ε
D	Е	Α	В	С

Casualizing the lines (2, 4, 5, 1, 3)

### Casualization in the Latin Square Design



Casualizing the columns (3, 5, 1, 4, 2)

### Casualization in the Latin Square Design

• 3x3 Latin Square



Randomize all but the first row

С	Α	В
В	С	Α
Α	В	С

# Advantages of the Latin Square Design

• You can control variation in two directions

Hopefully you increase efficiency as compared to the RCBD



## Disadvantages of Latin Square Design

- The number of treatments must equal the number of replicates
- The experimental error to increase with the size of the square
- Small squares have very few degrees of freedom for experimental error

## Disadvantages of Latin Square Design

• Effect of the Size of the Square on Error Degrees of Freedom

SOV	Df	2x2	3x3	4x4	5x5	8x8
Rows	r-1	1	2	3	4	7
Columns	r-1	1	2	3	4	7
Treatments	r-1	1	2	3	4	7
Error	(r-1)(r-2)	0	2	6	12	42
Total	$r^2 - 1$	3	8	15	24	63

The experimental error is likely to increase with the size of the square

### Statistical Model

• The statistical model for a latin square design is written as:

$$y_{ijk} = \mu + \tau_i + \gamma_i + \beta_j + \gamma_k + \varepsilon_{ijk}$$

where  $y_{ijk}$  is the observed response,  $\mu$  is the overall mean,  $\tau_i$  is the treatment effect,  $\beta_i$  is the row effect,  $\gamma_k$  is the column effect and  $\varepsilon_{ijk}$  is the experimental error

## Statistical Model

- One or both of the blocking factors can be treated as random (Mixed Models)
- Treatment effects are usually considered fixed

## Statistical Model

If rows and columns are random, we assume that:

We assume that:

- $\beta_j \sim N(0, \sigma_\beta^2)$
- $\gamma_k \sim N(0, \sigma_{\gamma}^2)$
- $e_{ijk} \sim N(0, \sigma_e^2)$
- $\beta_{j}$ ,  $\gamma_k$  and  $e_{ijk}$  are all independent of each other

## Analysis of Variance

Source of Variation	Degrees of Freedom	Sum of Squares	Mean Square
Rows	t-1	SS <sub>Row</sub>	$MS_{Row} = \frac{SS_{Row}}{t-1}$
Columns	t-1	SS <sub>Col</sub>	$MS_{Col} = \frac{SS_{Col}}{t-1}$
Treatments	t-1	SS <sub>Trt</sub>	$MS_{Trt} = \frac{SS_{Trt}}{t-1}$
Residuals	(t-1)(t-2)	$SS_{Res}$	$MS_{Res} = \frac{SS_{Res}}{(t-1)(t-2)}$
Total	$t^2 - 1$	$SS_Total$	

- There are (t-1)(t-2) degrees of freedom for the residual
  - More treatment levels increase the residual degrees of freedom, but require a larger experiment

## Analysis of Variance

Source of Variation	Degrees of Freedom	Mean Square	Expected Mean Square	F-Statistic
Rows	t - 1	$MS_{Row}$	$\sigma_{\varepsilon}^2 + t \sigma_{\beta}^2$	
Columns	t-1	$MS_{Col}$	$\sigma_{\varepsilon}^2 + t \sigma_{\gamma}^2$	
Treatments	t-1	$MS_{Trt}$	$\sigma_{\varepsilon}^2 + \frac{t}{t-1}\sum_i \tau_i^2$	$F = \frac{\mathrm{MS}_{\mathrm{Trt}}}{\mathrm{MS}_{\mathrm{Res}}}$
Residuals	(t-1)(t-2)	$MS_{Res}$	$\sigma_{\varepsilon}^2$	
Total	$t^2 - 1$	-		

The F-test for treatment effects is not affected if rows and/or columns are fixed

 The first exercise data set consists of stem dry weight (in g, log10 scale) of different sunflower genotypes



 These data come from Povin, C (1993) ANOVA: Experiments in Controlled Environments. Design and Analysis of Ecological Experiments. Ed. SM Scheiner, Ed. J Gurevitch. Chapman & Hall. 46-67.



- Use the R function read.csv to import the data
- Fit the model with fixed (and/or random) effects
- Check if model assumptions are met
- Build the ANOVA table and test the null hypothesis of no difference between genotypes
- Use multiple comparisons to assess pairwise differences









• Fit the model with fixed (and/or random) effects

```
#Fixed Model
fm <- lm(log_dry_weight ~ row + column + genotype, data = dados)
anova(fm)</pre>
```

```
Analysis of Variance Table
```

Response:	ിറ്റ	g_dry_	weig	ht							
	Df	Sum	Sq	Mean Sq	F value	Pr(	>F)				
row	5	0.730	072 C	).146145	4.2691	0.0083	606 *	<b>*</b> *			
column	5	1.256	<u>90 0</u>	).251380	7.3431	0.0004	727 *	**			
genotype	5	0.579	70 0	).115939	3.3867	0.0223	142 *	ł			
Residuals	20	0.684	67 0	034233.							
Signif. co	odes	s: 0	• * * *	'' 0.001	'**' 0.(	01'*'	0.05	•••	0.1	، ،	1

"fm" used to refer to an adjusted model

- Fit the model with fixed (and/or random) effects
  - Only genotype as fixed

```
library(lmerTest)
fme <- lmer(log_dry_weight ~ genotype + (1 | row) + (1 | column), data = dados)
anova(fme)</pre>
```

> fme <- lmer(log\_dry\_weight ~ genotype + (1 | row) + (1 | column), data = dados)
> anova(fme)
Type III Analysis of Variance Table with Satterthwaite's method
 Sum Sq Mean Sq NumDF DenDF F value Pr(>F)
genotype 0.5797 0.11594 5 20 3.3867 0.02231 \*
--Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

The package "ImerTest" is used to perform statistical hypothesis testing on linear mixed models fitted with the Ime4 package

Multiple Comparisons

```
library(emmeans)
fm_means <- emmeans(fm, "genotype")
pairs(fm_means)
plot(fm_means, comparisons = TRUE)
cld(fm_means, adjust = "tukey", Letters = letters)</pre>
```

 The emmeans function, short for "Estimated Marginal Means": used to calculate and extract estimated marginal means after fitting a statistical model

Multiple Comparisons

```
> library(emmeans)
> fm_means <- emmeans(fm, "genotype")</pre>
> pairs(fm_means)
                    SE df t.ratio p.value
contrast estimate
A - B
        -0.1290 0.107 20 -1.208
                                 0.8280
       0.1979 0.107 20 1.853
                                  0.4570
A - C
A - D
      0.0568 0.107 20 0.531
                                  0.9942
      -0.0418 0.107 20 -0.391
A - E
                                  0.9986
A - F
      0.2282 0.107 20 2.136
                                  0.3095
B - C
       0.3270 0.107 20 3.061
                                  0.0591
B - D
       0.1858 0.107 20 1.739
                                  0.5235
B - E
      0.0873 0.107 20 0.817
                                  0.9609
B - F
       0.3572 0.107 20 3.344
                                  0.0330
C - D
          -0.1412 0.107 20 -1.321
                                  0.7702
         -0.2397 0.107 20 -2.244
C - E
                                  0.2622
C - F
      0.0302 0.107 20
                          0.283
                                  0.9997
D - E
         -0.0985 0.107 20 -0.922
                                  0.9361
D - F
       0.1714 0.107 20
                         1.604
                                  0.6052
          0.2699 0.107 20
                           2.527
                                  0.1632
E - F
```

Results are averaged over the levels of: row, column P value adjustment: tukey method for comparing a family of 6 estimates

Multiple Comparisons



## References

- Chapters 9.10, 9.11 (for a more classical view)<sup>1</sup>
- Section 3.6.3 and sections 5.1 to 5.3<sup>2</sup>

1. Steel, R. G. & Torrie, J. H. Principles and Procedures of Statistics: A Biometrical Approach. 2nd Edition. (1980).

2. Casella, G. Statistical Design. (2008).