PRO 5971 - Statistical Process Monitoring

How to measure the performance of a control chart

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Outline

RUN LENGTH

The parameters of a control chart: sample size n, the control limits (UCL and LCL) and sampling interval (for simplicity, let us consider h = 1)

Action: whenever the statistic is plotted out of the control limits, a search for special causes starts.

Performance measure: the number of samples until a signal. In Figure 1, 4 samples for the first signal; 6 samples for second signal; 6 samples for third signal, etc



Figure 1: # of samples until a signal

Run length

Y= # of samples until a signal - is a random variable known as RUN LENGTH If the monitored statistics are **independent** \rightarrow Y follows a Geometric distribution with parameter *p*, the probability to signal

Its probability function is

$$P(Y = y) = p(1 - p)^{y-1}, y = 1, 2, ...$$

For example: $\begin{cases} P(Y = 1) = p, & \text{the probability to signal at the first sample} \\ P(Y = 2) = p(1 - p), & \text{the probability to signal at the second sample} \\ \dots P(Y = 5) = p(1 - p)^4, & \text{the probability to signal at the fifth sample} \end{cases}$

 $E(Y) = \frac{1}{p}$ is the average number of samples until a signal.

It is known as AVERAGE RUN LENGTH (ARL) and has been the most used a performance metric in SPC.

Other measures: median run length (MRL); standard deviation run length(SDRL).

When the process is stable (in control or under
$$H_0$$
)
$$\begin{cases} \alpha - \text{probability to signal} \\ 1 - \alpha, \text{ probability to not signal} \\ ARL_0 = \frac{1}{\frac{1}{\text{probability to signal}}} = \frac{1}{\alpha} \end{cases}$$

 $ARL_0 \begin{cases} \text{is the average number of sample until a signal when the process is in-control} \\ \text{It is desirable to have larger values for } ARL_0 \\ \text{Usually values like 370 or 500 are used as } ARL_0 \text{ in manufacturing process monitoring} \\ \text{IMPORTANT: The value of } ARL_0 \text{ must have a practical context meaning} \end{cases}$

When the process is out of control (or under H_1) $\begin{cases} 1 - \beta - \text{probability to signal} \\ \beta, \text{ probability to not signal} \\ ARL_1 = \frac{1}{\text{probability to signal}} = \frac{1}{1 - \beta} \end{cases}$

 $ARL_1 \begin{cases} \text{is the average number of sample until a signal when the process is out-of-control} \\ \text{It is desirable to have } ARL_1 \approx 1 \end{cases}$

Plan \overline{X} control charts (assuming σ known and remains stable) to have $ARL_0 = 200, 370, 500$ with n = 3, 5, 9. Obtain ARL_1 when μ_0 shifts to $\mu_1 = \mu_0 \pm \delta\sigma$ for $\delta = 0.25, 0.5, 1, 1.5, 2, 2.5$.

Plan S^2 control charts (assuming that μ is stable) to have $ARL_0 = 200,370,500$ with n = 10, 15, 25. Obtain ARL_1 when μ_0 shifts to $\sigma_1^2 = \sigma_0^2 \delta$ for $\delta = 0.25, 0.5, 1, 1.5, 2, 2.5$.

Plan *p* control charts to have $ARL_0 = 200, 370, 500$ with n = 20, 30, 50. Obtain ARL_1 when $p_0 = 0.05, 0.1$ shifts to $p_1 = p_0\delta$ for $\delta = 0.25, 0.5, 1, 1.5, 2, 2.5$. Use the exact and the approximate distribution of the monitored statistic.

Plan *u* control charts to have $ARL_0 = 200, 370, 500$ with n = 10, 20, 30. Obtain ARL_1 when $\lambda_0 = 0.05, 5$ shifts to $\lambda_1 = \lambda_0 \delta$ for $\delta = 0.25, 0.5, 1, 1.5, 2, 2.5$. Use the exact and the approximate distribution of the monitored statistic.

In each control chart discuss the behavior of ARL_1 . Is it possible to find an optimal design for that control chart?

When the monitored statistics are not independent, the determination of ARL is more complicated.

There are many alternatives for its obtaining.

Three alternatives will be presented, two based on Markov chains and the third by simulation.

A simple example of Markov chain

Consider a consumer support center. The aim is to discover averagely how many calls a customer needs to make in order to be attended. The process can be summarized as the following steps:

- 1 The customer makes a call due to some reason.
- 2 The call may be completed or not.

3 - If completed and some attendant is available, the customer is finally attended. Otherwise, the system leaves a message to call latter.



Figure 2: Consumer support

cont - A simple example

Markov chain - the current event (state) depends on the previous one. The consumer support process can be described by the states:

- State A the customer makes a call
- State B the call is completed
- State C the attendant answers

The whole process can be represented by a transition matrix of probability

	A	В	С
А	рдд	P AB	PAC
В	р _{ВА}	р _{ВВ}	рвс
С	РСА	Р СВ	рсс

The index ij in matrix denotes: from state i to state j

 p_{ij} - denotes probability to reach the current state j conditioned that at previous step was at state i - p(j|i).

For example: AB - from state A to state B; p_{AB} - probability to the call be completed conditioned(B) that the customer makes a call (A) - p(B|A).

 p_{BC} - probability to the attendant answers (state c) conditioned that the call is completed (state B) - p(C|B)

Assuming $p_{AB} = p_{BC} = 0.5$, the transition matrix of probability is

		A	В	С
P =	А	0.5	0.5	0
	В	0.5	0	0.5
	С	0	0	1

The states A and B are transitory and C is absorbent

 ${\bf P}$ - the full transitory matrix

 ${\bf Q}$ -a square sub matrix of ${\bf P}$ after the removal of rows and columns of absorbent states.

 $N = (I - Q)^{-1}$

 $ARL = (\mathbf{I} - \mathbf{Q})^{-1}\mathbf{1}$

In our example, we have

$$\mathbf{N} = \begin{bmatrix} 4 & 2 \\ 2 & 2 \end{bmatrix}$$

That is $n_{11}=4$, averagely 4 calls are made before be completed; $n_{21}=2$, even completed (state B), averagely more two calls are needed.

$$\mathbf{ARL} = \begin{bmatrix} 4 & 2 \\ 2 & 2 \end{bmatrix} \times \begin{bmatrix} 1 \\ 1 \end{bmatrix} = \begin{bmatrix} 6 \\ 4 \end{bmatrix}$$

 $ARL_0=6$

Details see Brook & Evans (1972)

Method 2 - ARL by Markov chain - renewal process

The same states can be considered, but the main idea is that whenever the state C is reached, the process is renewed (or feed forward), that is, a new customer will make a call and reaches the state A or B. The new transitory matrix P^* becomes

$$\mathbf{P}^* = \begin{array}{c|cccc} A & B & C \\ \hline A & 0.5 & 0.5 & 0 \\ B & 0.5 & 0 & 0.5 \\ C & 0.5 & 0.5 & 0 \end{array}$$

As $s \to \infty$, the rows of \mathbf{P}^{*s} are equal to a vector $\boldsymbol{\pi} = [\pi_A, \pi_B, \pi_c]$.

Each π_i measures the probability to reach the state *i* for a long term.

The vector $\pi = [\pi_A, \pi_B, \pi_c]$ can be obtained as solution of the system of equations $\pi = \pi \mathbf{P}^*$ such that $\pi_A + \pi_B + \pi_c = 1$

In our example, $\pi = [3/6, 2/6, 1/6]$, thus the average run length is

$$rac{1}{Prob ext{ of state C}} = rac{1}{(1/6)} = 6$$

When the process is stable, the non-conformance rate is 0.2. A sample of five items X_1, X_2, \ldots, X_5 from a Poisson (0.2) is taken. Thus $Y = \sum_{i=1}^{5} X_i$ follows Poisson (1).

The engineer desires to monitor the average rate. For that, he proposed the criterion: if $Y_i > 4$ or if $2 \le Y_{i-1} \le 4$ and $2 \le Y_i \le 4$ then the process is stopped to search the special causes.

Define the states of the Markov chain for this process monitoring. Obtain the ARL0.

State A, if $0 \leq Y \leq 1$

State B, if $2 \le Y \le 4$

State C, the process is declared out-of-control

Matrix of transitory probabilities

$$\mathbf{P}^{*} = \begin{array}{c|cccc} A & B & C \\ \hline A & p_{A} & p_{B} & 1 - p_{A} - p_{B} \\ B & p_{A} & 0 & 1 - p_{A} \\ C & p_{A} & p_{B} & 1 - p_{A} - p_{B} \end{array}$$

The stationary distribution is

$$\pi = \left[p_A; rac{p_B}{1+p_B}; rac{1-p_A-p_Ap_B}{1+p_B}
ight]$$

$$ARL_0 = \frac{1 + p_B}{1 - p_A - p_A p_B}$$

It is known that Shewhart control chart has good performance to detect large shifts.

So to improve these charts, several contributors have added/included new supplementary rules in the decision criterion.

Let consider an example of supplementary rule for the \overline{x} control chart.

Two sets of limits: control and warning are used: LCL and UCL are respectively the lower ad upper control limits and LWL and UWL, respectively the lower and upper warning limits with LCL < LWL < UWL < UCL.

[LWL,UWL] is known as the central region; $[\textit{LCL},\textit{LWL}] \cup [\textit{UWL},\textit{UCL}]$ is the warning region.

Decision rules to declare the process as out-of-control:

Rule 1: $\overline{X}_i > UCL$;

Rule 2: $\overline{X}_i < LCL$;

Rule 3: at least two points of the last 3 monitored statistics in the warning region [*LCL*, *LWL*]; that is, *LCL* $< \overline{X}_i < LWL$ and *LCL* $< \overline{X}_{i-1} < LWL$ or *LCL* $< \overline{X}_i < LWL$, *LCL* $< \overline{X}_{i-2} < LWL$ and *LWL* $< \overline{X}_{i-1} < UWL$

Rule 4: at least two points in the last 3 observations in the warning region [UWL, UCL] that is, $UWL < \overline{X_i} < UCL$ and $UWL < \overline{X_{i-1}} < UCL$ or $UWL < \overline{X_i} < UCL$, $UWL < \overline{X_{i-2}} < UCL$ and $LWL < \overline{X_{i-1}} < UWL$

For simplicity let us consider only rules which the monitored statistics fell or above the center line or below the center line, that is, or R1 and R3 or R2 and R4.

The probabilities of a monitored statistic fell in the central region, warning region and above the control limit are respectively: a_1 , a_2 and $1 - a_1 - a_2$.

Define the states of Markov chain and determine an expression for ARL

States: rules 2 and 4

- A last two observations in the central region
- A1 the last but one observation in the warning region and the last in central region

B - the last but one observation in the central region and the last in warning region

 ${\sf C}$ - the current observation in the action region; the last two observations in the warning region; the last but one in the central region but the third to last and the last observations in the warning region

The transitory matrix

	A	A1	В	С
А	a_1	0	a 2	1- <i>a</i> 1- <i>a</i> 2
A1	a_1	0	0	$1-a_1$
В	0	a_1	0	$1-a_1$
С	0	0	0	1

cont - Solution - Exercise - ARL by Markov chain - not easy - with absorbent state

$$(\mathbf{I} - \mathbf{Q})^{-1} = \frac{1}{d} \begin{bmatrix} 1 & a_1 a_2 & a_2 \\ a_1 & 1 - a_1 & a_1 a_2 \\ a_1^2 & (1 - a_1) a_1 & 1 - a_1 \end{bmatrix}$$
$$d = 1 - a_1 - a_1^2 a_2$$
$$(\mathbf{I} - \mathbf{Q})^{-1} \mathbf{1} = \begin{bmatrix} \frac{1 + a_1 a_2 + a_2}{1 - a_1 - a_1^2 a_2} \\ \frac{1 - a_1 - a_1^2 a_2}{1 - a_1 - a_1^2 a_2} \\ \frac{1}{1 - a_1 - a_1^2 a_2} \end{bmatrix}$$

As always we assume that the process start at state A then ARL= $\frac{1 + a_1a_2 + a_2}{1 - a_1 - a_1^2a_2}$ ARL₀ if a_1 and a_2 are obtained under μ_0 ;

ARL₁ if a_1 and a_2 are obtained under μ_1 ;

cont - Solution - Exercise - ARL by Markov chain - not easy - with absorbent state

The overall
$$ARL_0=0.5 \frac{1+a_1a_2+a_2}{1-a_1-a_1^2a_2}$$
 by symmetry.
 $\mu_1 = \mu_0 + \delta\sigma$, thus ARL_1 can be expressed as $ARL_1(\delta)$ for the rules 2 and 4
Thus we have $ARL_1(-\delta)$ for the rules 1 and 3,
The overall

$$ARL_1 = \left[rac{1}{ARL_1(-\delta)} + rac{1}{ARL_1(\delta)}
ight]^{-1}$$

The same states in the previous slides but once state C reached, the process is renewed The transitory matrix becomes

	A	A1	В	С
А	a_1	0	a 2	1- <i>a</i> 1- <i>a</i> 2
A1	a_1	0	0	$1-a_1$
В	0	a_1	0	$1-a_1$
С	a ₁	0	a_2	$1 - a_1 - a_2$

Finding its stationary distribution

$$\pi = \begin{bmatrix} \pi_1 = \frac{a_1(a_1+a_2+1)}{1+a_2+a_1+a_2} & \pi_2 = \frac{(a_1+a_2)}{1+a_2+a_1+a_2} & \pi_3 = \frac{a_2}{1+a_2+a_1+a_2} & \pi_4 = \frac{1-a_1^2a_1-a_1}{1+a_2+a_1+a_2} \end{bmatrix}$$

So $ARL = \frac{1}{\pi_4} = \frac{1+a_2+a_1+a_2}{1-a_1^2a_1-a_1}$

Some reasons to use the simulation to determine the average run length and also the control (warning) limits:

- the distribution of the monitored statistic is unknown
- the determination of control and warning limits of the control chart are quite complicated or do not have closed forms
- to check the analytical results.

Two runs of simulations are needed:

1 - simulated the monitored statistics when the process is in-control.

Usually for the determination of control and warning limits such that meet the target $\ensuremath{\textit{ARL}}_0$ value

2 - simulated the monitored statistics when the process is out-of-control: To calculate $\ensuremath{\textit{ARL}}_1.$

A simplified algorithm for in-control process:

- 1. Define the decision rules to judge the process as out-of-control
- 2. Choose the initial control and warning limits, etc,
- 3. Simulate a large number of the monitored statistics (under in-control parameters)
- Count the number of the monitored statistics until satisfied the decision rules. This is the first in-control run length RL₁, save this count.
- 5. Repeat the steps 3-4 several times until to get a large number of in-control run lengths (like 10 thousand)
- 6. Calculate its average as $\widehat{ARL}_0 = \frac{\sum_{i=0}^{1000} RL_i}{10000}$ and verify if it is closer to the target ARL_0 value. If it is closer, stop the simulation, you have determined the control limits and use them in the second run of simulation. Otherwise choose other control (warning) limits for step 2 and repeat steps 3-6.

A simplified algorithm for out-of-control process:

- 1. Use the outputs of the first run (in-control simulation) as control and warning limits, etc,
- 2. Simulate a large number of the monitored statistics (under out-of-control parameters)
- Count the number of the monitored statistics until satisfied the decision rules. This is the first out-of-control run length *RL*₁; save this count.
- 4. Repeat steps 2-3 several times until to get a large number of out-of-control run lengths (like 10 thousand)

5. Calculate its average as $\widehat{ARL}_1 = \frac{\sum_{i=1}^{10000} RL_i}{10000}$. This an estimate of ARL_1

1 - Choose control and warning limits and other necessary in-control parameters (as in-control mean and variance and sample size) for "the not easy problem". Develop Monte Carlo simulations to check the theoretical results of ARL_0 by Markov Chain.

2 - Choose shift sizes for the process mean and check by Monte Carlo simulation if the expression stated for ARL_1 is correct.

3 - Now consider that some target value of $ARL_0 = \tau$ is desirable. Search for the control and warning limits by simulation such that minimize ARL_1 for a shift of one standard deviation.

Define the states of a Markov chain associated to these problems. Find their stationary distribution and interpret their results.

- 1. Assume that the chance of a rain tomorrow only depends on today's8 weather condition. If today rains, tomorrow will rain with probability α ; if today does not rain, tomorrow will rain with probability β .
- 2. A communication system transmits using the codes 0 or 1. Each digit transmitted must pass through several stages. At n-th stage, *p* is the probability of not change the code transmitted in the system communication.
- 3. John's humor can be classified into 3 classes: bad-tempered (A), more or less tempered (B) or good-tempered (C). If today he is at state A, the probabilities of being at states A, B, and C tomorrow are respectively 0.5; 0.4 and 0.1. If today he is at state B, the probabilities of being at states A, B, and C tomorrow are respectively 0.3; 0.4 and 0.3. And if he is at state C, the probabilities of being at states A, B, and C tomorrow are respectively 0.2; 0.3 and 0.5.

Exercises Markov Chain II

- 4. The possibility of raining or not today depends on the climate conditions of last two days. If the last 2 days has rained, tomorrow will rain with probability of 0.7; if yesterday not rained but today rained, then tomorrow will rain with probability 0.5; if yesterday rained yesterday and today not rained, them tomorrow will rain with probability 0.4. If last 2 days has not rained, then tomorrow will rain with probability 0.2.
- 5. An insurance company applies a bonus system to their customers according to their frequency of occurrences (Assume that the frequency of the occurrences follows a Poisson distribution of parameter $\lambda = 0.5$). The value of their insurance of the next year is a function of these occurrences. The clients are classified into 4 classes according to this table:

		Class of the next year			
		# of the occurrences			
Current class	Value of insurance	0	1	2	≥ 3
A	200	А	В	С	D
В	250	А	С	D	D
С	400	В	D	D	D
D	600	С	D	D	D

- 6. John practices running every morning. He may leave by the front door or back door (with equal probability) and after the running he may come in by the front door or back door (also with equal probability). He has five pairs of tennis and as soon as he gets in, he takes off them at the entrance no matter which door he comes in. If no pair of tennis is in entrance when he leaves in the morning, he runs barefoot.
- 7. Consider the following decision rule of a process monitoring: a sample size of n is split into two sub-samples of size n_1 and $n_2 = n n_1$, with $n_1 > n_2$. Each n_1 item is classified according to some criteria as approved or not. Let Y, be the number of not approved items. If $Y \leq CL_y$, the process is declared in-control; otherwise, measure the n_2 items and calculate its average. If $\overline{X} > CL_x$ then the process is declared out-of-control, the process is adjusted and a new sample of size n is taken and the described procedure continues.

- 8. The following rule has been used to monitor a process: step 1) take a sample of size n_a and classify each one as approved or not according to a rule A. Let Y_a be the number of not approved items. If Y_a ≤ CL_a then the process is declared in-control and go to step 2; step 2) take a sample of size n_b < n_a and classify each item as approved or not according to criteria B. Let Y_b, the number of items disapproved. If Y_b ≤ CL_b then the process is judged as in-control, otherwise is said to be out-of-control. step 3: Repeat the step 1.
- 9. Khoo's scheme: whenever $\overline{X} > UCL$ the process is declared out-of-control or a sequence of size 2 values of \overline{X} in the range [UWL; UCL].
- Extend the problem 9 considering two sided case, that is: whenever X > UCL or X < LCL the process is declared out-of-control or a sequence of size 2 values of X in the range [UWL; UCL] or in the range [LCL; LWL].

References

Brook, D. & Evans, D. (1972), 'An approach to the probability distribution of cusum run length', *Biometrika* pp. 539–549.